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## Ordered Risk Aggregation under Dependence Uncertainty

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Risk measure: Let  $\mathcal{M}$  be the set of cdfs on  $\mathbb{R}$ . A risk measure is defined as

 $\rho: \mathcal{M} \to \mathbb{R}.$ 

For  $F \in \mathcal{M}$ , if  $X \sim F$ , we also write  $\rho(X) = \rho(F)$ .

Examples: For  $F \in \mathcal{M}$ ,

Left Value-at-Risk:

 $\operatorname{VaR}_q^L(F) = F^{-1}(q) = \inf\{t \in \mathbb{R} : F(t) \ge q\}, q \in (0, 1].$ 

• Right Value-at-Risk:

$$\operatorname{VaR}_{p}^{R}(F) = F^{-1}(p+) = \inf\{t \in \mathbb{R} : F(t) > p\}, p \in [0, 1)$$

Expected Shortfall:

$$\mathrm{ES}_p(F) = \frac{1}{1-p} \int_p^1 \mathrm{VaR}_u^R(F) \mathrm{d} u, p \in (0,1).$$

Range-VaR:

$$\operatorname{RVaR}_{p,q}(F) = \frac{1}{q-p} \int_{p}^{q} \operatorname{VaR}_{u}^{R}(F) \mathrm{d}u, 0 \leq p < q < 1.$$

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Depende	nce uncertainty			

Suppose that we have a portfolio of two risks  $\boldsymbol{X}$  and  $\boldsymbol{Y}.$  We are interested in

 $\rho(X+Y).$ 

We assume:

- Known mariginal distributions;
- Uncertain dependence structure;
- Order constraint:  $X \leq Y$  almost surely.

For  $F, G \in \mathcal{M}$  such that F is stochastically smaller than G i.e.,  $F \ge G$ , define the set

$$\mathcal{F}_2^o(F,G) = \{(X,Y) : X \sim F, Y \sim G, X \leqslant Y\}.$$

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Risk boi	unds with order o	constraint		

The worst-case and best-case values of  $\rho(X + Y)$  over the set  $\mathcal{F}_2^o(F, G)$  denoted by

$$\overline{\rho}(\mathcal{F}_2^o(F,G)) := \sup\{\rho(X+Y) : (X,Y) \in \mathcal{F}_2^o(F,G)\},\$$

and

$$\underline{\rho}(\mathcal{F}_2^o(F,G)) := \inf\{\rho(X+Y) : (X,Y) \in \mathcal{F}_2^o(F,G)\}.$$

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Concave	order			
Concave				

A distribution F is called smaller than a distribution G in *concave order*, denoted by  $F \leq_{cv} G$ , if  $\int \phi \, \mathrm{d}F \leq \int \phi \, \mathrm{d}G$  for all concave  $\phi : \mathbb{R} \to \mathbb{R}$ , provided that both integrals exist.

For a risk measure  $\rho:\mathcal{M}\to\mathbb{R},$  we define three commonly used properties:

- A risk measure  $\rho$  is monotone if  $\rho(F) \leq \rho(G)$  whenever  $F \leq_{st} G$ ;
- A risk measure ρ is ≤<sub>cv</sub>-consistent if ρ(F) ≤ ρ(G) whenever F ≤<sub>cv</sub> G;
- A risk measure ρ is ≤<sub>cx</sub>-consistent if ρ(F) ≤ ρ(G) whenever G ≤<sub>cv</sub> F.

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The concave ordering bounds of X + Y?

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### Concave ordering in unconstrained case

Concave ordering bounds of X + Y:

- A random vector (X, Y) is *comonotonic* if there exists a random variable U and two increasing functions f and g such that X = f(U) and Y = g(U) almost surely.
- A random vector (X, Y) is *countermonotonic* if (X, -Y) is comonotonic.
- It is well known that

$$X^{c} + Y^{c} \leq_{\mathrm{cv}} X + Y \leq_{\mathrm{cv}} X^{co} + Y^{co}$$

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Countermonotonicity may violate the order constraint!

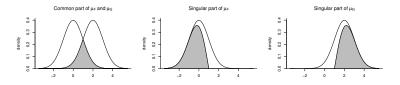
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DL cou	oling			

Directional Lower (DL) coupling: Arnold et al. (2020) and Nutz and Wang (2020)

Denote by  $\mu_F$  and  $\mu_G$  the Borel probability measures generated by continuous distributions F and G, respectively.

- The common part μ<sub>F</sub> ∧ μ<sub>G</sub> of F and G is the maximal measure θ such that θ ≤ μ<sub>F</sub> and θ ≤ μ<sub>G</sub>.
- The singular parts of F and G are defined as  $\mu'_F = \mu_F \mu_F \wedge \mu_G$  and  $\mu'_G = \mu_G \mu_F \wedge \mu_G$ .

Figure: F = N(0, 1) and G = N(2, 1)



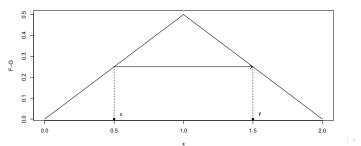
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DL cou	oling			

The DL coupling between F and G has two parts.

- The common part of F and G couples identically with each other.
- The transport from the singular part of *F* to the singular part of *G*, denoted by  $T^{F,G}$ , is defined as

$$T^{F,G}(x) = \inf \left\{ z \geqslant x : F(z) - G(z) < F(x) - G(x) \right\}.$$

Figure: 
$$F = \text{Unif}[0, 2]$$
 and  $G = \text{Unif}[1, 2]$ 



F-G

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### Concave orderings in constrained case

#### Lemma

For  $(X, Y), (X^c, Y^c), (X', Y') \in \mathcal{F}_2^o(F, G)$  such that  $(X^c, Y^c)$  is comonotonic and (X', Y') is DL-coupled, we have

$$X^{c} + Y^{c} \leqslant_{\mathrm{cv}} X + Y \leqslant_{\mathrm{cv}} X' + Y'.$$

#### Corollary

Suppose that  $(X, Y), (X^c, Y^c), (X', Y') \in \mathcal{F}_2^o(F, G)$  such that  $(X^c, Y^c)$  is comonotonic and (X', Y') is DL-coupled. If  $\rho$  is  $\leq_{cv}$ -consistent, then

$$\underline{\rho}(\mathcal{F}_2^o(F,G)) = \rho(X^c + Y^c) \leqslant \rho(X + Y) \leqslant \rho(X' + Y') = \overline{\rho}(\mathcal{F}_2^o(F,G)).$$

If  $\rho$  is  $\leq_{cx}$ -consistent, then

$$\underline{\rho}(\mathcal{F}_2^o(F,G)) = \rho(X'+Y') \leqslant \rho(X+Y) \leqslant \rho(X^c+Y^c) = \overline{\rho}(\mathcal{F}_2^o(F,G)).$$

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Tail risk	measures			

Let  $F^{[p,1]}$  be the upper *p*-tail distribution of  $F \in \mathcal{M}$ , namely

$${\mathcal F}^{[p,1]}(x)=rac{({\mathcal F}(x)-p)_+}{1-p},\quad x\in{\mathbb R}.$$

#### Definition (Liu and Wang (2020))

For  $p \in (0, 1)$ , a risk measure  $\rho$  is a *p*-tail risk measure if  $\rho(F) = \rho(G)$  for all  $F, G \in \mathcal{M}$  such that  $F^{[p,1]} = G^{[p,1]}$ .

For a *p*-tail risk measure  $\rho$ , there always exists another risk measure  $\rho^*$ , called the generator, such that  $\rho(F) = \rho^* \left(F^{[p,1]}\right)$ . We call  $(\rho, \rho^*)$  a *p*-tail pair of risk measures.

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<i>p</i> -conce	ntration			

*p*-concentration:

A *p*-tail event of a random variable X is an event A ∈ A with
0 < P(A) = 1 − p < 1 such that X(ω) ≥ X(ω') holds for all ω ∈ A and ω' ∈ A<sup>c</sup>.

A random vector (X, Y) is *p*-concentrated if X and Y shares a common *p*-tail event of probability 1 - *p*.

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### Bounds on tail risk measures

DL coupling:  $(X, Y) \sim D_*^{F,G}$ .

### Theorem

Suppose that  $F \leq_{st} G$ ,  $p \in (0, 1)$ ,  $(\rho, \rho^*)$  is a p-tail pair of risk measure, and  $\rho^*$  is monotone and  $\leq_{cv}$ -consistent. We have

$$\overline{\rho}(\mathcal{F}_2^o(F,G)) = \overline{\rho^*}\left(\mathcal{F}_2^o\left(F^{[p,1]},G^{[p,1]}\right)\right) = \rho^*(X+Y),$$

where  $(X, Y) \sim D_*^{F^{[p,1]}, G^{[p,1]}}$ .

The class of  $\leq_{cv}$ -consistent generators  $\rho^*$ :

•  $\rho^* = \text{ess-inf}$ , corresponding to  $\rho = \text{VaR}_p^R$ ;

• 
$$\rho^* = \mathbb{E}$$
, corresponding to  $\rho = \mathrm{ES}_{\rho}$ ;

• 
$$\rho^* : X \mapsto -\text{ES}_t(-X)$$
, corresponding  $\rho = \text{RVaR}_{\rho,q}$ , where  $t = (1-q)/(1-p)$ .

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### VaR bounds

#### Proposition

For continuous distributions F and G such that F  $\leqslant_{\rm st}$  G and  $p\in(0,1),$  we have

$$\overline{\operatorname{VaR}}_{p}^{R}(\mathcal{F}_{2}^{o}(F,G)) = \min\left\{\inf_{x \in [F^{-1}(p+1), G^{-1}(p+1)]} \left\{T^{F^{[p,1]}, G^{[p,1]}}(x) + x\right\}, 2G^{-1}(p+1)\right\},\$$

and

$$\underline{\operatorname{VaR}}_{p}^{L}(\mathcal{F}_{2}^{o}(F,G)) = \max\left\{\sup_{x \in [F^{-1}(p), G^{-1}(p)]} \left\{x + \hat{T}^{F^{[0,p]}, G^{[0,p]}}(x)\right\}, 2F^{-1}(p)\right\}$$

where  $\hat{T}^{F^{[0,p]},G^{[0,p]}}(x) = \sup \left\{ t \leqslant x : F^{[0,p]}(t) - G^{[0,p]}(t) < F^{[0,p]}(x) - G^{[0,p]}(x) \right\}.$ 

Unconstrained problem: Rüschendorf (1982)

#### Proposition

Suppose that F and G are strictly increasing continuous distribution functions such that F  $\leqslant_{\rm st}$  G. For  $p\in(0,1),$  we have

 $\overline{\operatorname{VaR}}^L_p(\mathcal{F}^o_2(F,G)) = \overline{\operatorname{VaR}}^R_p(\mathcal{F}^o_2(F,G)) \quad \text{ and } \quad \underline{\operatorname{VaR}}^L_p(\mathcal{F}^o_2(F,G)) = \underline{\operatorname{VaR}}^R_p(\mathcal{F}^o_2(F,G)).$ 

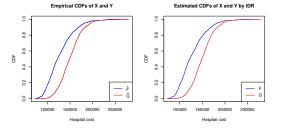
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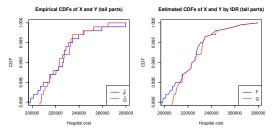
• The aggregate loss S = X + Y where  $X \sim F$  and  $Y \sim G$  represent the losses caused by females and males, respectively, from a portfolio of 50 males and 50 females

- $X \leq Y$  is reasonable due to many common risk factors
- Cannot reject the hypothesis  $\hat{\mathcal{F}}\leqslant_{\mathrm{st}}\hat{\mathcal{G}}$
- Estimate F and G such that F ≤<sub>st</sub> G using the isotonic distributional regression (Henzi et al. (2019))

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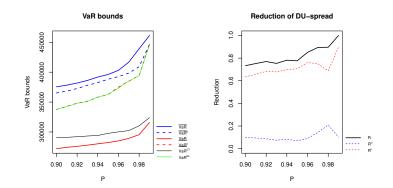
### Numerical results





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# Thank You!

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