

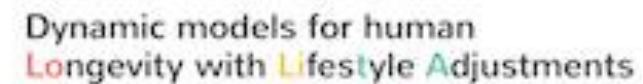
Creating risk indicators

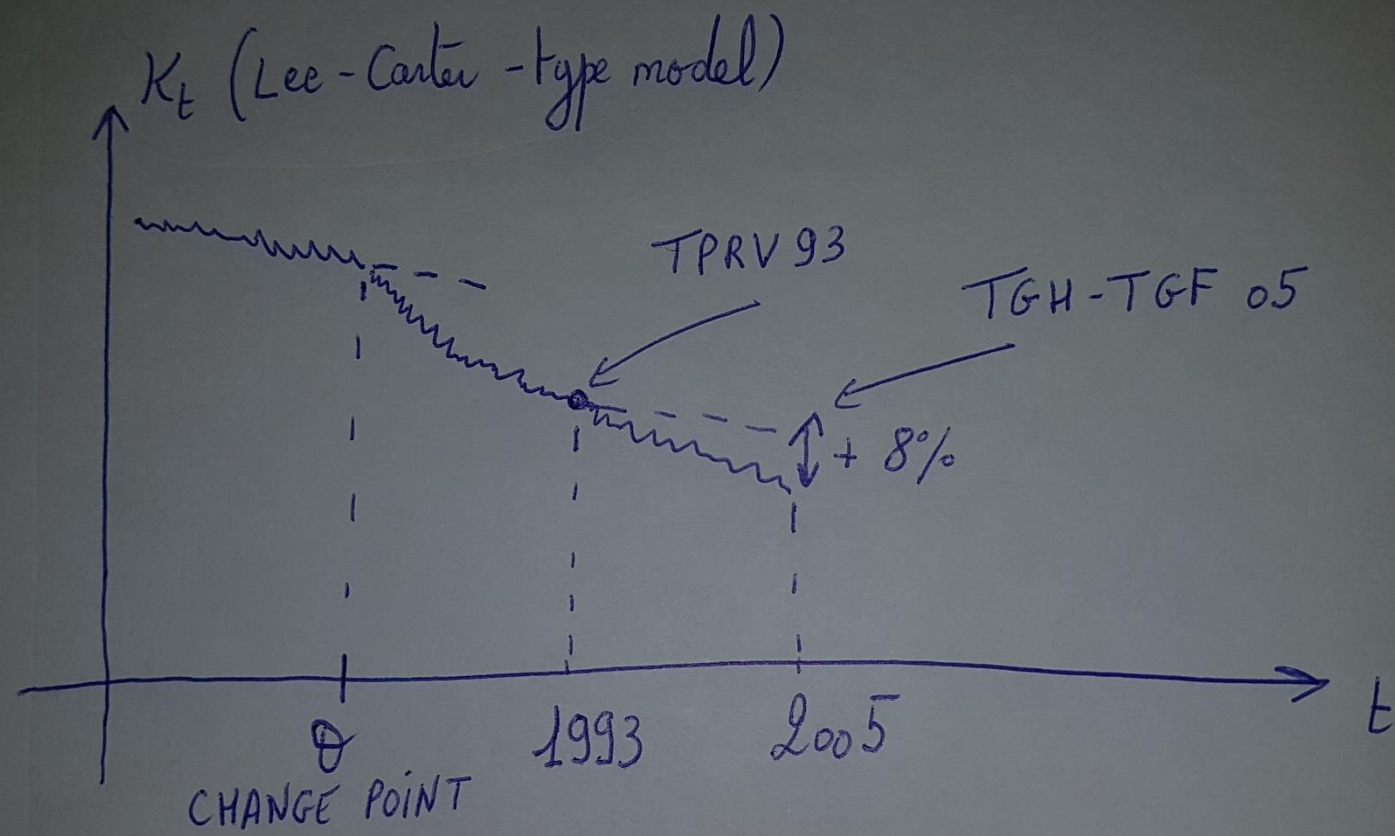
Stéphane Loisel

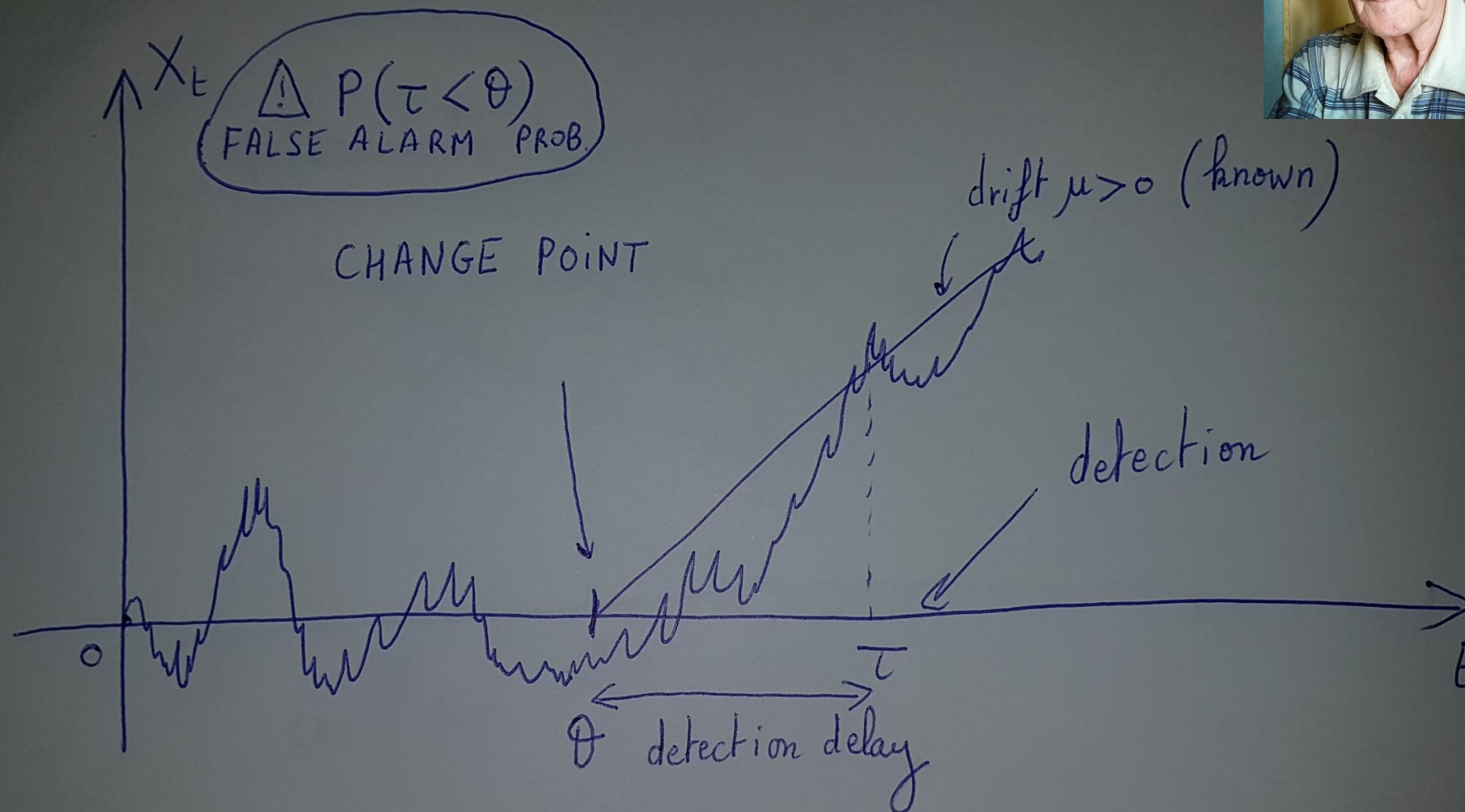
ISFA, Université Lyon 1

CERA Global Conference, 2021

Joint work with N. El Karoui and Y. Salhi









Bayesian setup for random change-point

Brownian framework with abrupt change in the drift

- ▶ Based on the conditional distribution of the time of change,
- ▶ Formulated as an optimal stopping problem
- ▶ Page(1954), Shiryaev(1963), Roberts(1966), Beibel(1988), Moustakides (2004), and many others...

Poisson framework with abrupt change in intensity

- ▶ Based on the conditional distribution of the time of change, with exponential or geometric prior distribution
- ▶ More recent studies : Gal (1971), Gapeev (2005), Bayraktar (2005, 2006), Dayanik (2006) for compound Poisson, Peskir, Shyriaev(2009) and others

MATHEMATICAL SETTINGS

We consider a portfolio of insured population:

- Let $N = (N_t)_{t \geq 0}$ be a **counting process** indicating the deaths of policyholders and $\lambda = (\lambda_t)_{t \geq 0}$ its **intensity**.
- The counting process N_t , is **available sequentially** through the filtration $\mathcal{F}_t = \sigma\{N_s, 0 < s \leq t\}$.
- We suppose that the insurance company relies on a **Cox-like** model to project her own experienced mortality:

$$\lambda_t = \underline{\rho} \lambda_t^0,$$

- λ_t^0 is a **reference intensity** and $\underline{\rho}$ is a positive parameter.
- λ^0 is considered deterministic and may refer whether to a projection of national population/best estimate...

Model risk/parameter uncertainty: **Change-point**

$$\lambda_t = \mathbf{1}_{\{t < \theta\}} \underline{\rho} \lambda_t^0 + \mathbf{1}_{\{t \geq \theta\}} \bar{\rho} \lambda_t^0.$$

Without loss of generality we can assume that $\underline{\rho} = 1$ and let $\rho = \bar{\rho} > 1$.

PROBABILISTIC FORMULATION

Let \mathbb{P}_θ (resp. $\mathbb{E}_\theta[\cdot]$) be the probability measure (resp. expectation) induced when the change takes place at time θ

Example

- For $\theta = 0$, the process is *out-of-control*
- For $\theta = \infty$, the process is *in-control*

Detect the change-point θ as quick as possible while avoiding false alarms

OPTIMALITY CRITERIA, LORDEN (1971)-LIKE

- The detection delay $\mathbb{E}_\theta \left[(N_\tau - N_\theta)^+ \middle| \mathcal{F}_\theta \right]$
- The frequency of false alarm $\mathbb{E}_\infty[N_\tau]$

OPTIMIZATION PROBLEM

OPTIMIZATION PROBLEM

Find τ^* such that $C(\tau^*) = \inf_{\tau} \sup_{\theta \in [0, \infty]} \text{ess sup } \mathbb{E}_{\theta} \left[(N_{\tau} - N_{\theta})^+ \middle| \mathcal{F}_{\theta} \right]$
subject to $\mathbb{E}_{\infty}[N_{\tau}] = \omega$.

ASSUMPTION

- 1 $\int_0^t \lambda_s ds < \infty, \quad \mathbb{P}_{\infty}, \mathbb{P}_0\text{-a.s.}$
- 2 $N_{\infty} = \infty \quad \mathbb{P}_{\infty}, \mathbb{P}_0\text{-a.s.}$

OPTIMALITY OF THE CUSUM PROCEDURE (1/7)

Let the Radon-Nikodym density of \mathbb{P}_0 with respect to \mathbb{P}_∞ be defined as

$$\frac{d\mathbb{P}_0}{d\mathbb{P}_\infty} \Big|_{\mathcal{F}_t} = \exp U_t,$$

where $U_t = \log(\rho)N_t + (1 - \rho) \int_0^t \lambda_s^0 ds$ is the log-likelihood ratio.

Let $V(x)$ be the CUSUM process; with head-start $0 \leq x < m$; defined as

$$V_t(x) = U_t - (-x) \wedge \underline{U}_t \tag{1}$$

where \underline{U}_t is the running infimum of U , i.e. $\underline{U}_t = \inf_{s \leq t} U_s$.

The process $V(x)$ measures the size of the drawup, comparing the present value of the process U to its historical infimum \underline{U} .

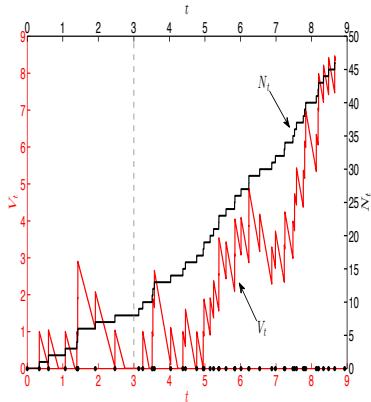
Let $\tau_m(x)$ be the first hitting time of $V(x)$ of the barrier m , i.e.

$$\tau_m(x) = \inf\{t \geq 0, V_t(x) \geq m\}.$$

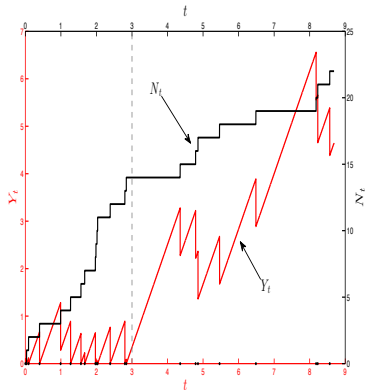
Theorem

If $\mathbb{E}_\infty[N_{\tau_m(0)}] = \omega$ then $\tau_m(0)$ is optimal, i.e. $\inf_\tau C(\tau) = C(\tau_m(0))$

Typical paths with change of regime at date 3

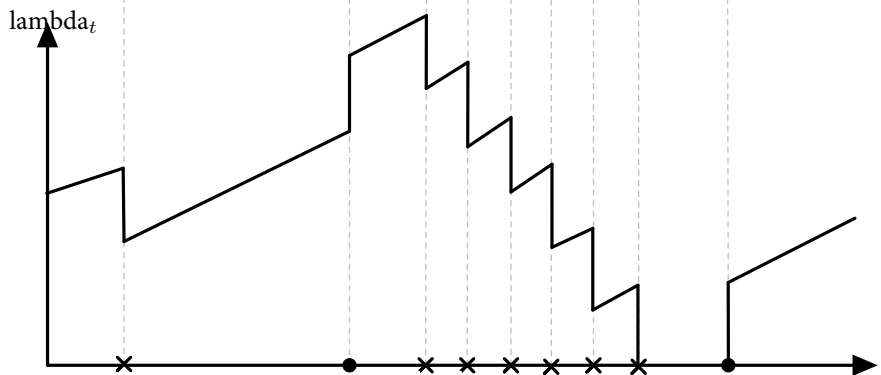
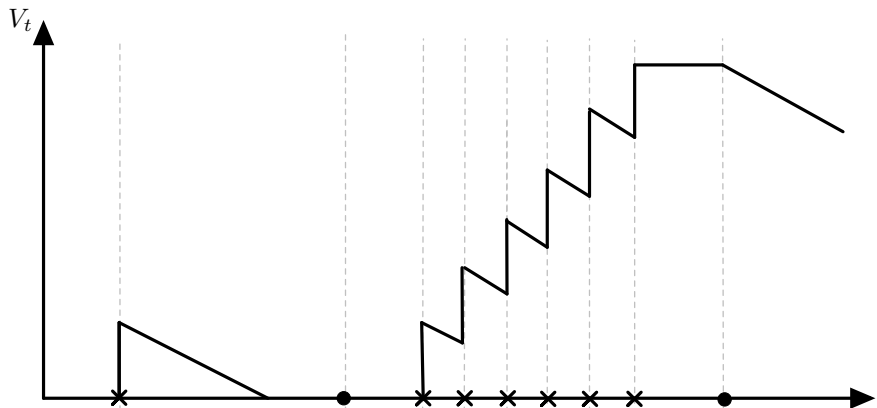


(a) Processes N and V_t



(b) Processes N and Y_t

Figure: Sample paths, for $\rho = 1.5$, of the cusum processes N , V^ρ (left) and N , Y_t^ρ for $\rho = 0.5$ (right).



Why is quick detection important in insurance?
How to choose parameter rho?

	females				males			
	Doubled improvements		Mortality level at 80% of the expected		Doubled improvements		Mortality level at 80% of the expected	
	pension value	interest rate	pension value	interest rate	pension value	interest rate	pension value	interest rate
55	+5.4%	+32bp	+3.1%	+19bp	+6.7%	+42bp	+3.7%	+24bp
65	+5.76%	+43bp	+4.7%	+36bp	+7%	+57bp	+5.7%	+48bp
75	+5.2%	+55bp	+7.6%	+80bp	+6.3%	+74bp	+9.1%	+107bp
85	+3.6%	+60bp	+13.2%	+207bp	+4.3%	+84bp	+15.4%	+281bp

TABLE: TGH05/TGF05 with flat interest rate of 3%

Monitoring Mortality

Sounding an alarm for the change $\rho^{\text{Hyp}} \rightarrow \rho^{\text{Target}}$

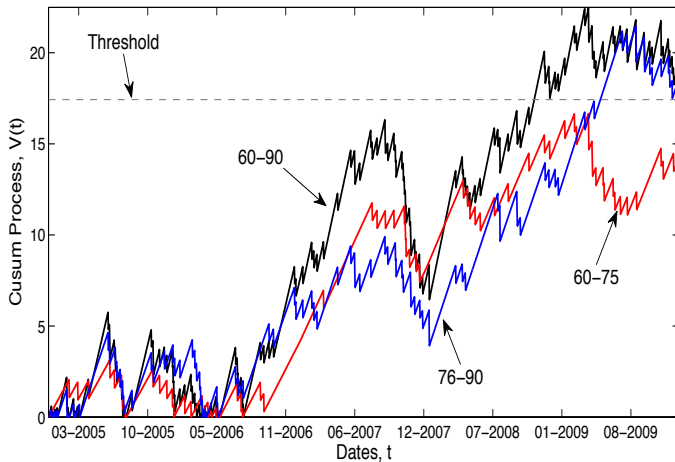
- We simulate deaths on the portfolio with different levels $\rho^{\text{Target}} = 95\%, 90\% \text{ and } 85\% \text{ s.t.}$

$$D(x, t) \sim \text{Pois}(\rho^{\text{Target}} \times L(x, t) \times \mu^{\text{ERM00}}(x, t))$$

- We suppose that *the actuary* made an assumption of $\rho^{\text{Hyp}} = 100\%$
- We set-up the monitoring/surveillance on the observed deaths and try to detect a change from $\rho^{\text{Hyp}} = 100\%$ to $\rho^{\text{Target}} = 95\%, 90\% \text{ and } 85\%$ respectively.
- We test different sizes of the portfolio small sized 1000, 5000 and a (relatively) large 10000 and compare the results

Monitoring Mortality

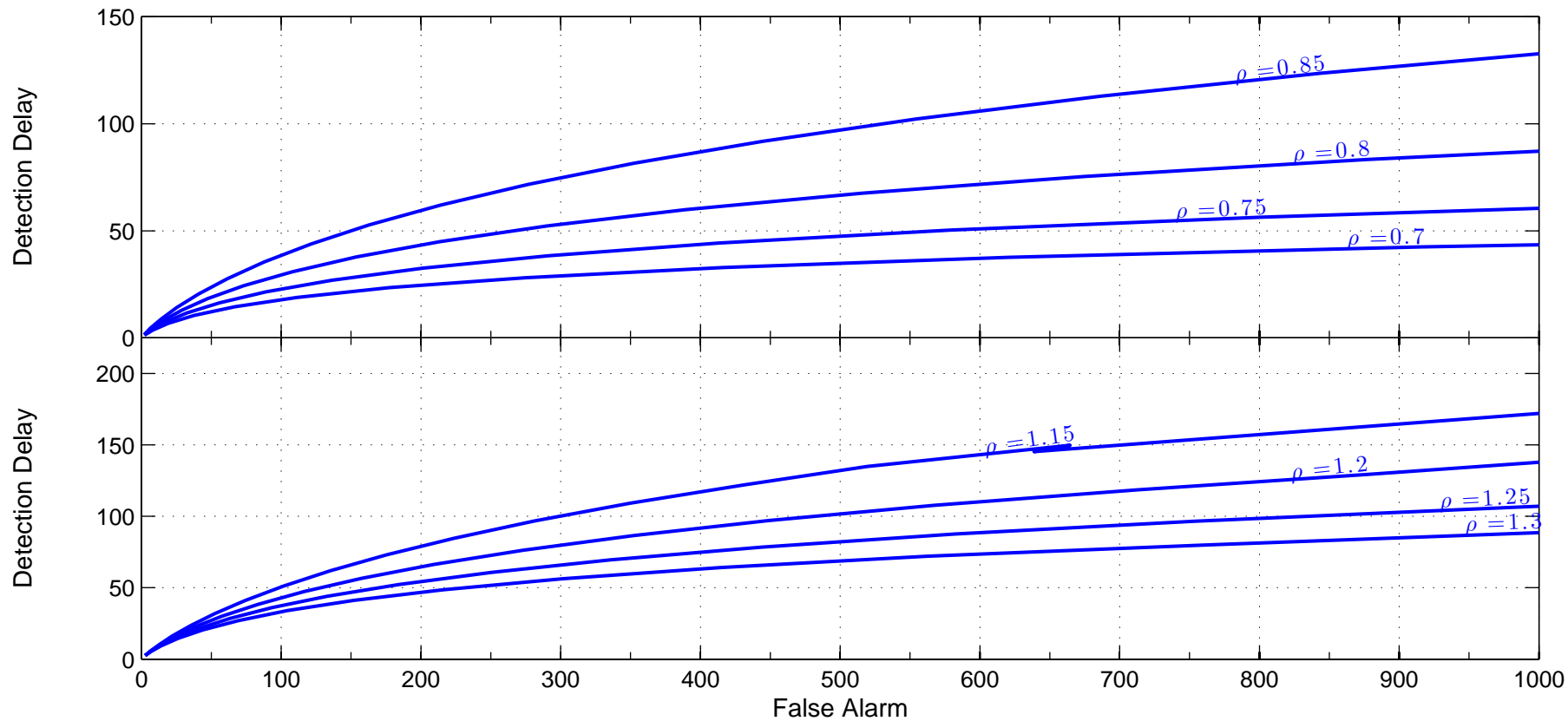
Sounding an alarm for the change $\rho^{\text{Hyp}} = 100\% \rightarrow \rho^{\text{Target}} = 95\%$

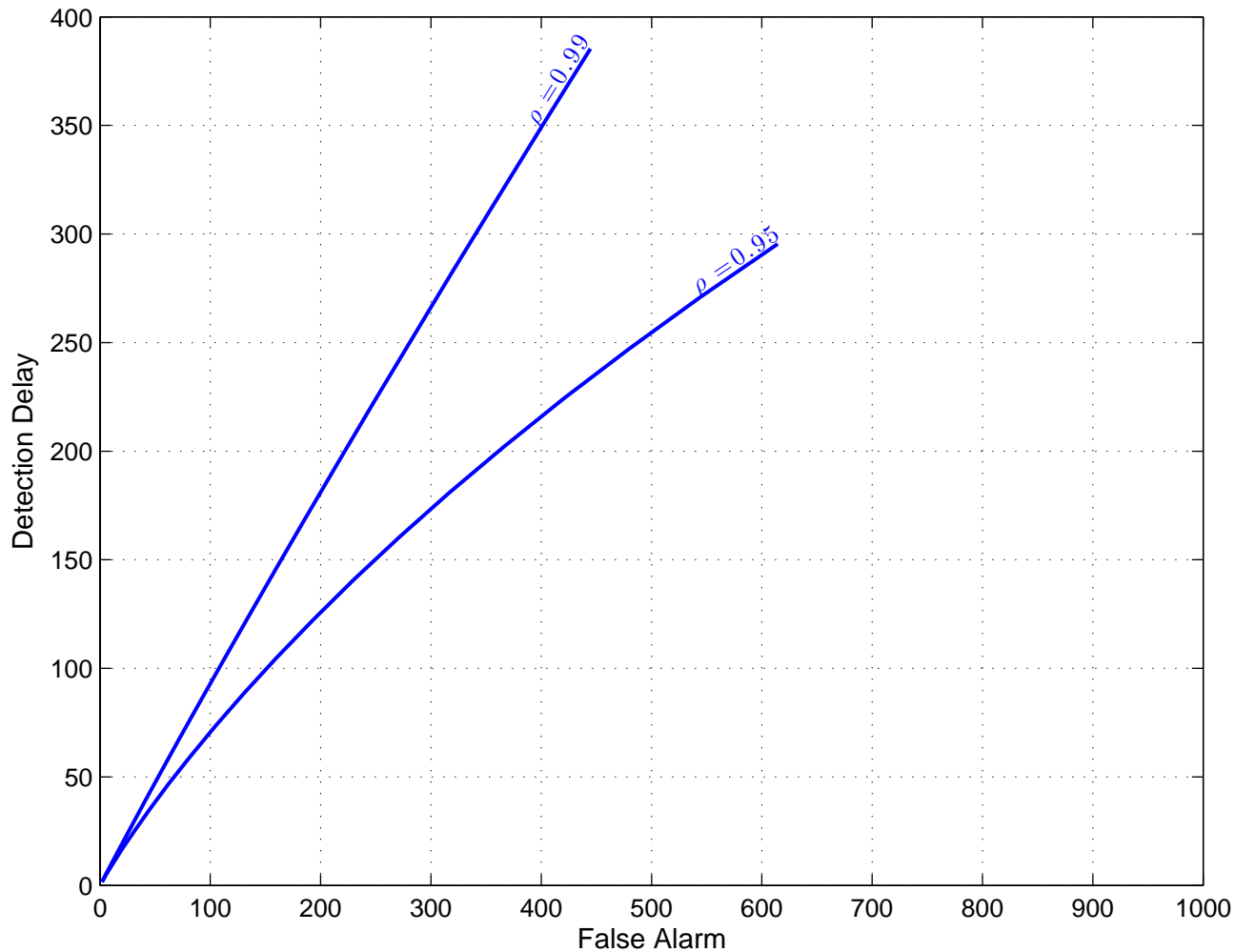


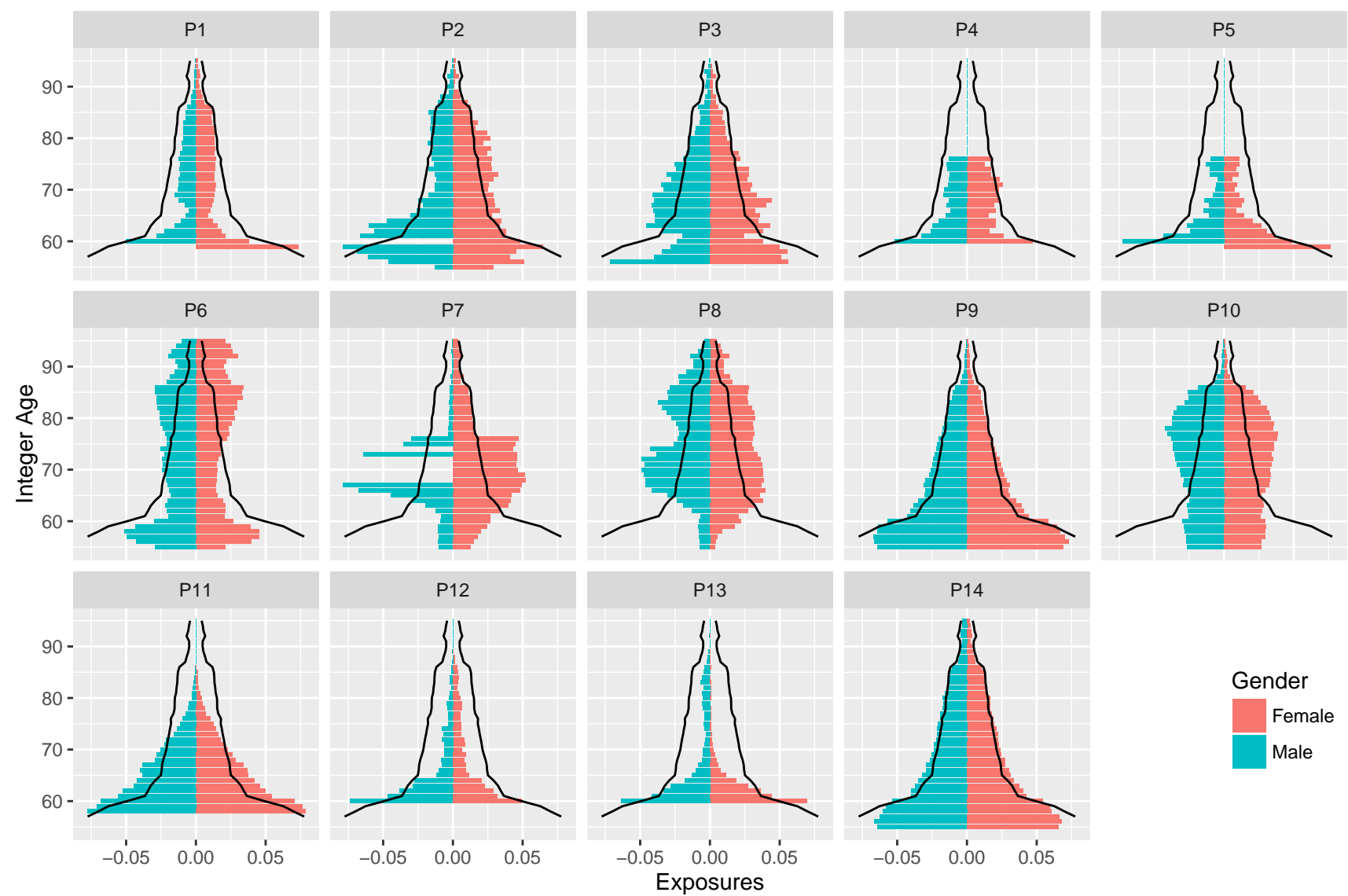
Detection Delay

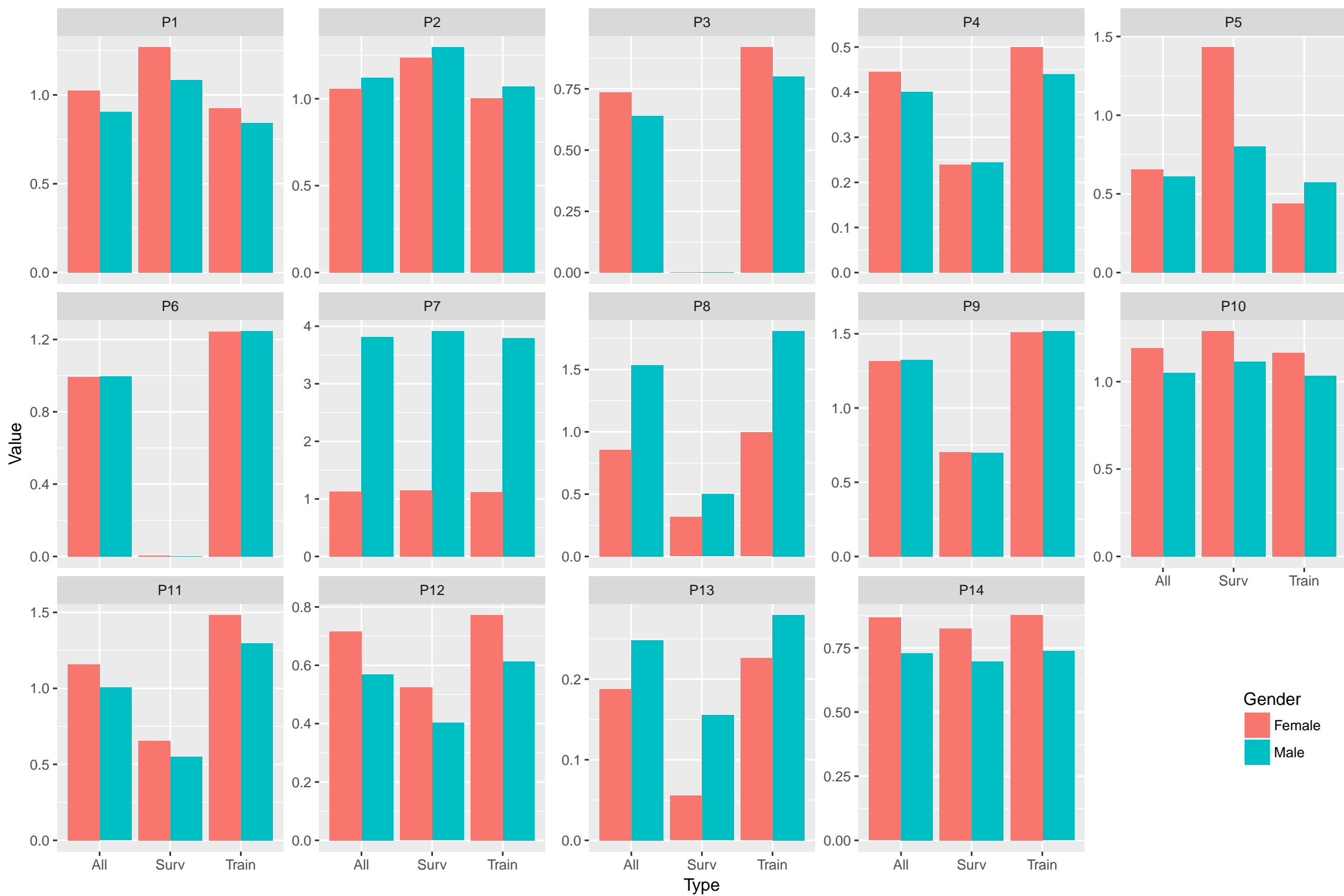
Impact of Portfolio Size and Age Tranches

Size		1000			5000			10000		
Hyp.	Ages	60-90	60-75	76-90	60-90	60-75	76-90	60-90	60-75	76-90
deaths	100% → 95%	596	710	498	246	99	107	240	99	106
	100% → 90%	244	320	186	106	55	59	112	55	58
	100% → 85%	92	122	100	58	35	36	61	34	36
time	100% → 95%	1086	1130	1120	576	617	422	308	327	212
	100% → 90%	931	1124	947	276	373	241	151	192	127
	100% → 85%	707	980	734	161	247	159	84	127	80

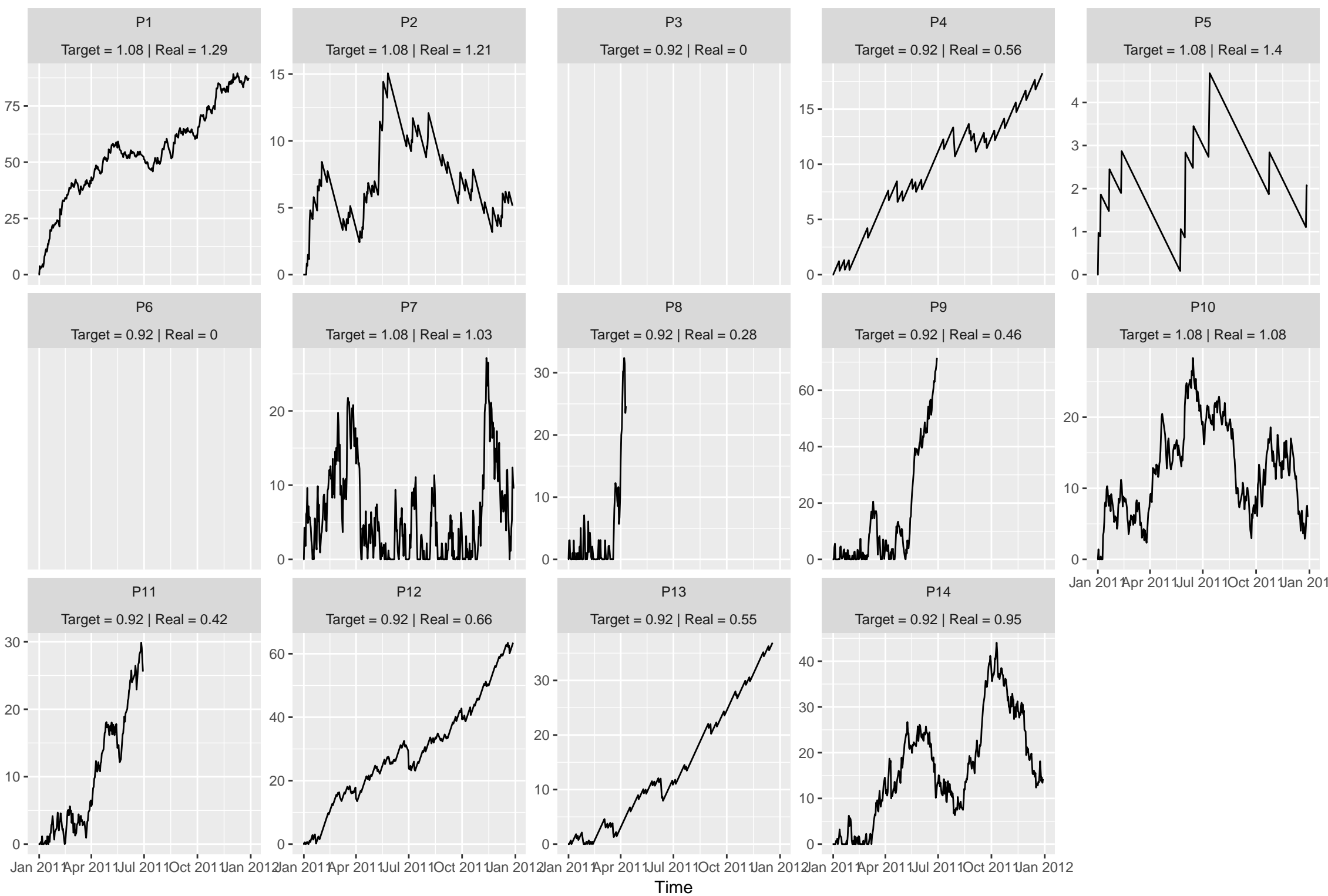




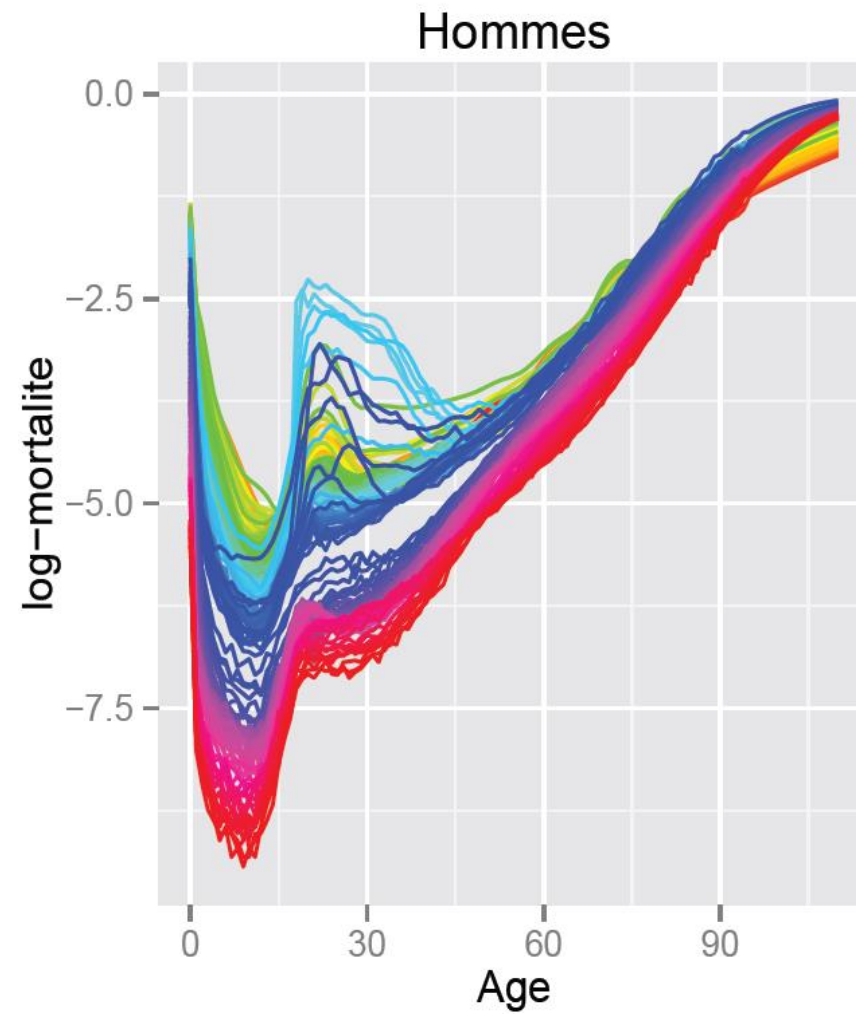




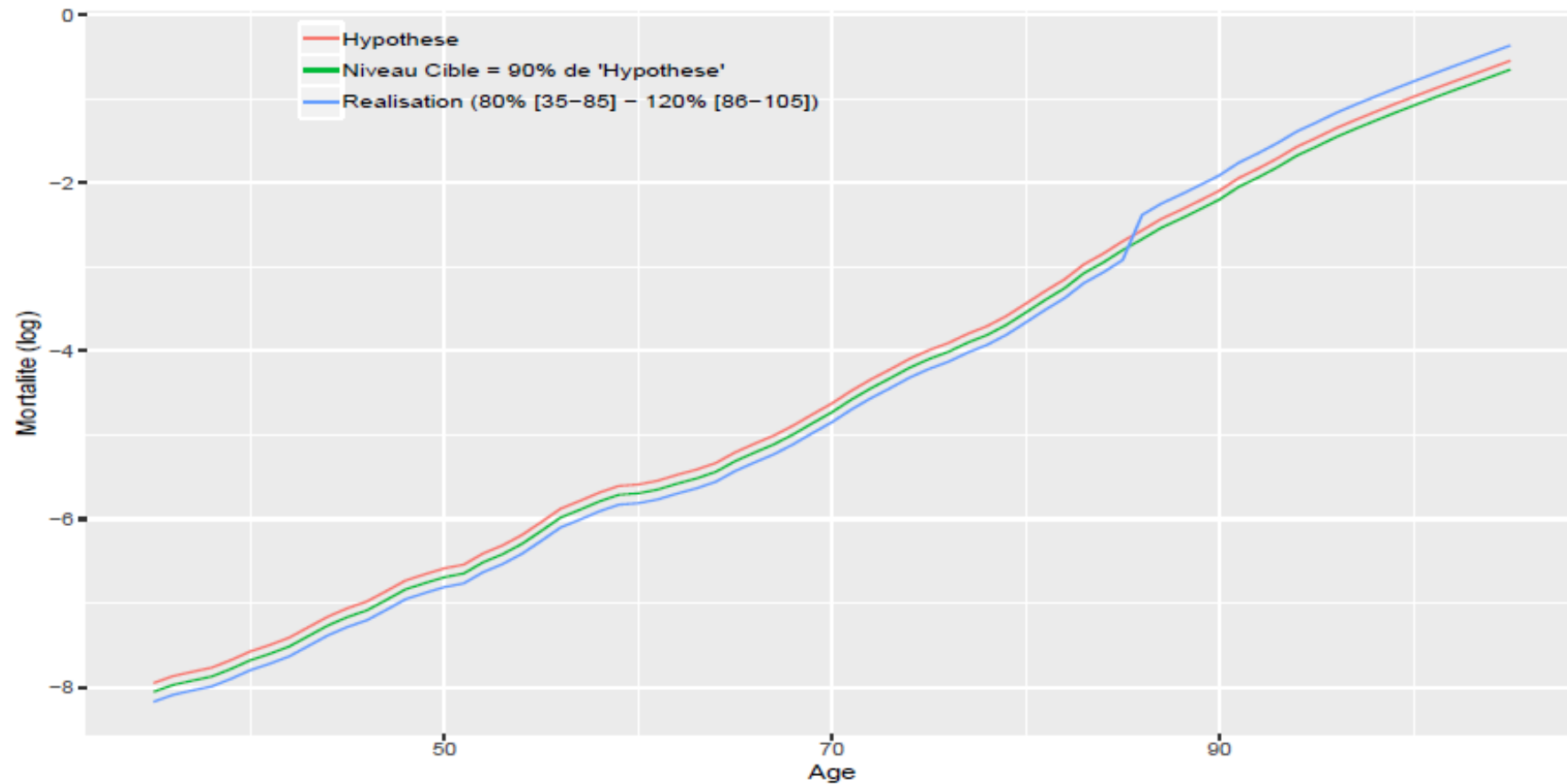
Cusum Process



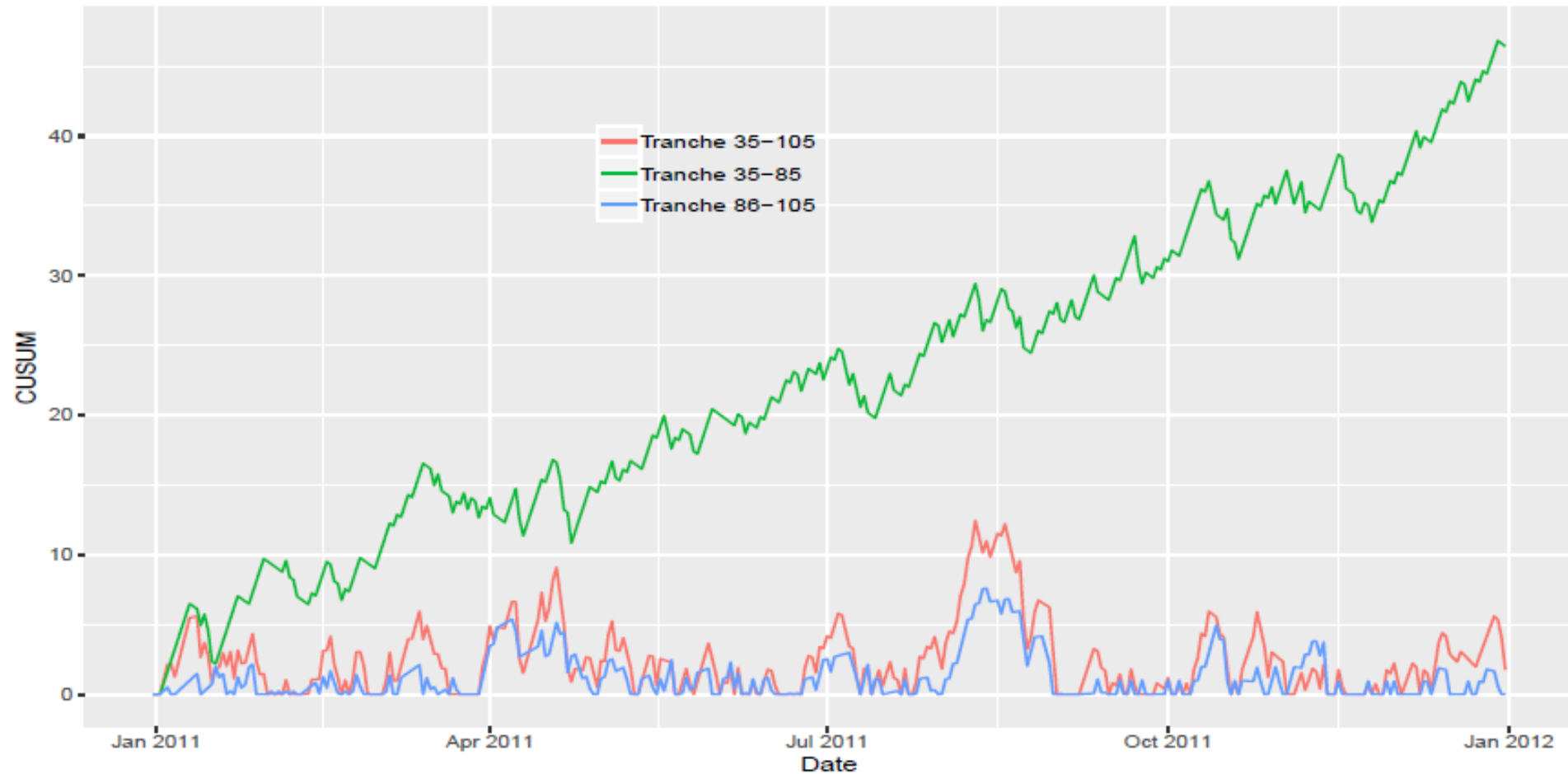
France, 1816-2000

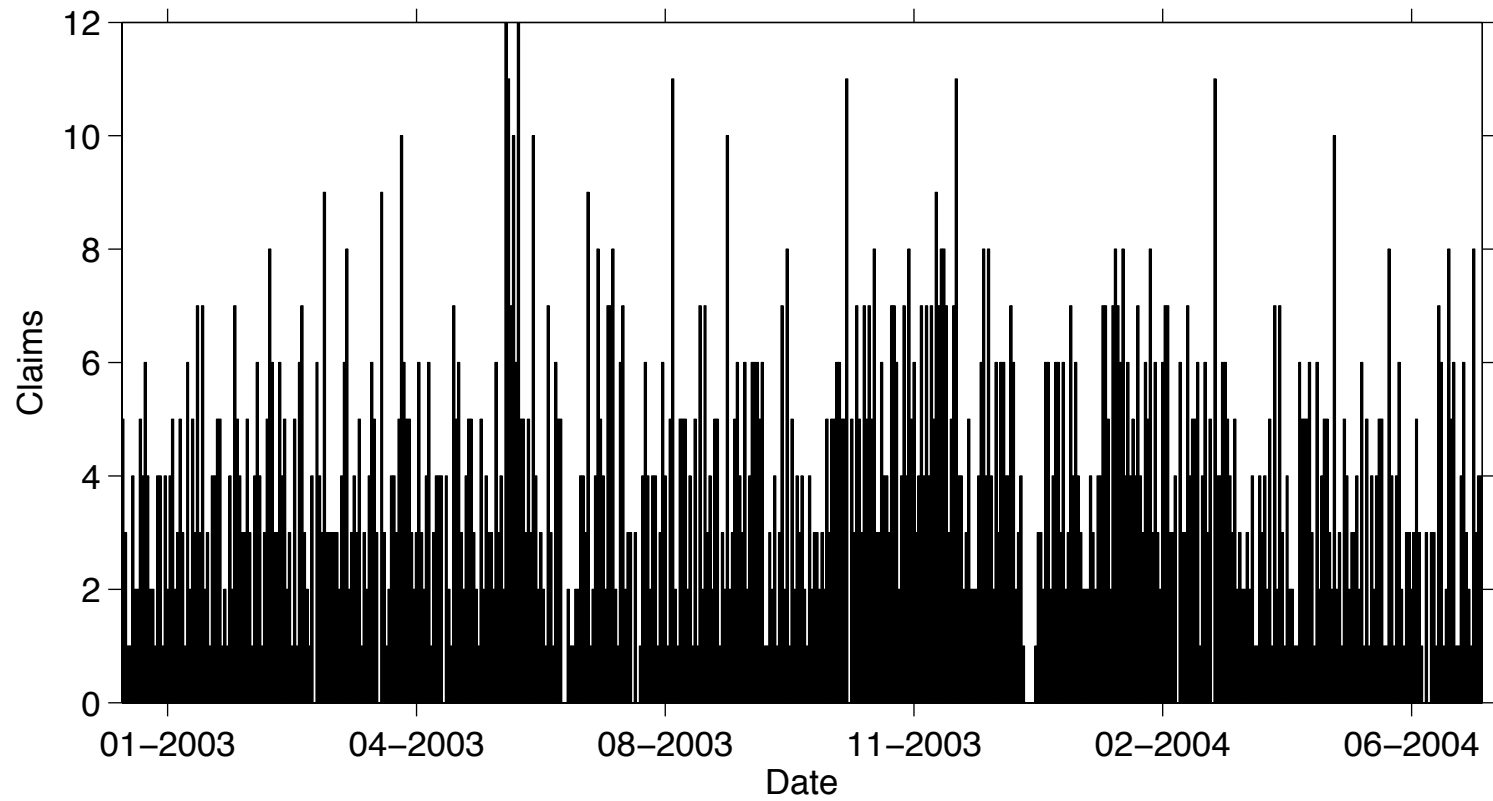


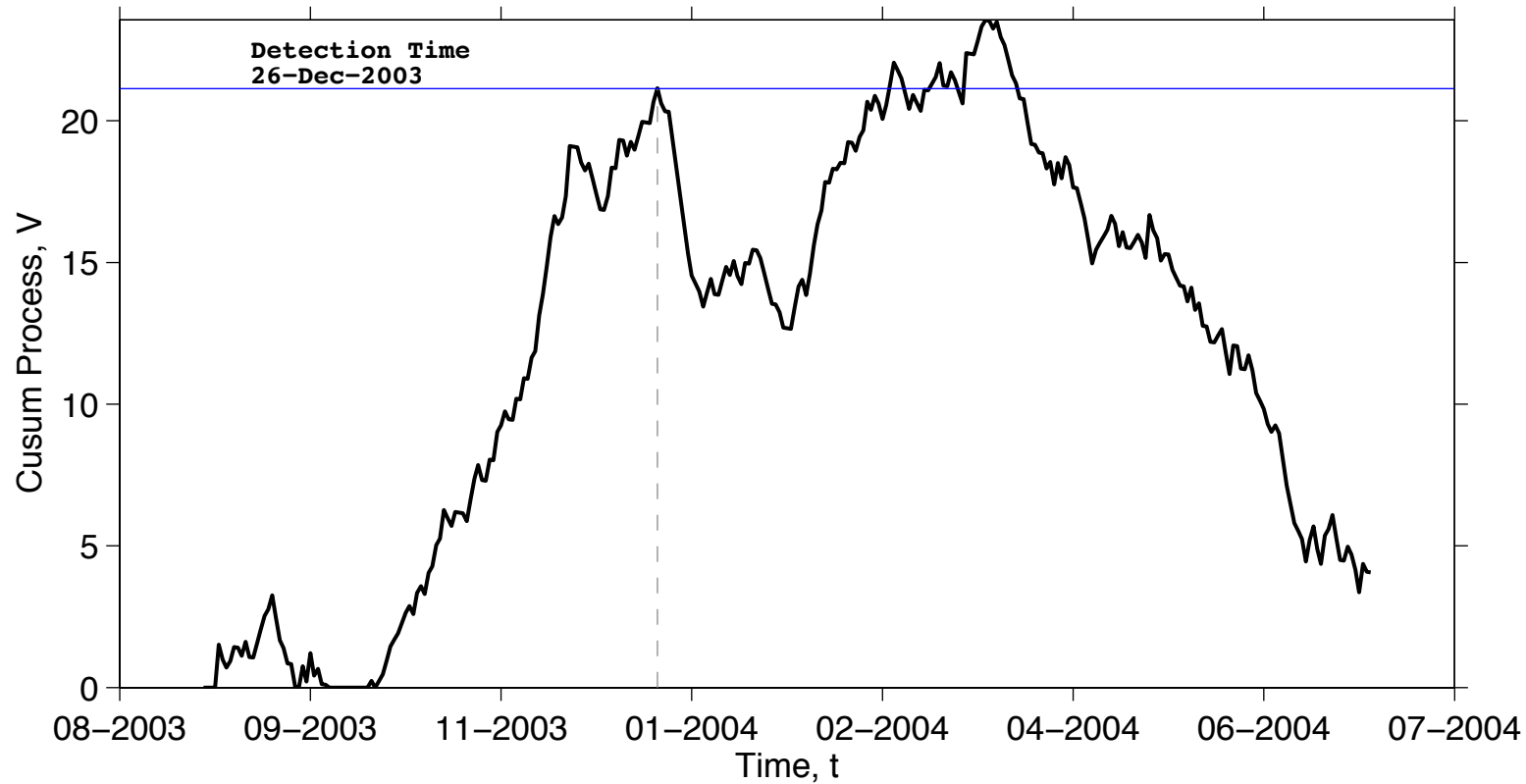
Impact of rectangularization (without transhumanism)

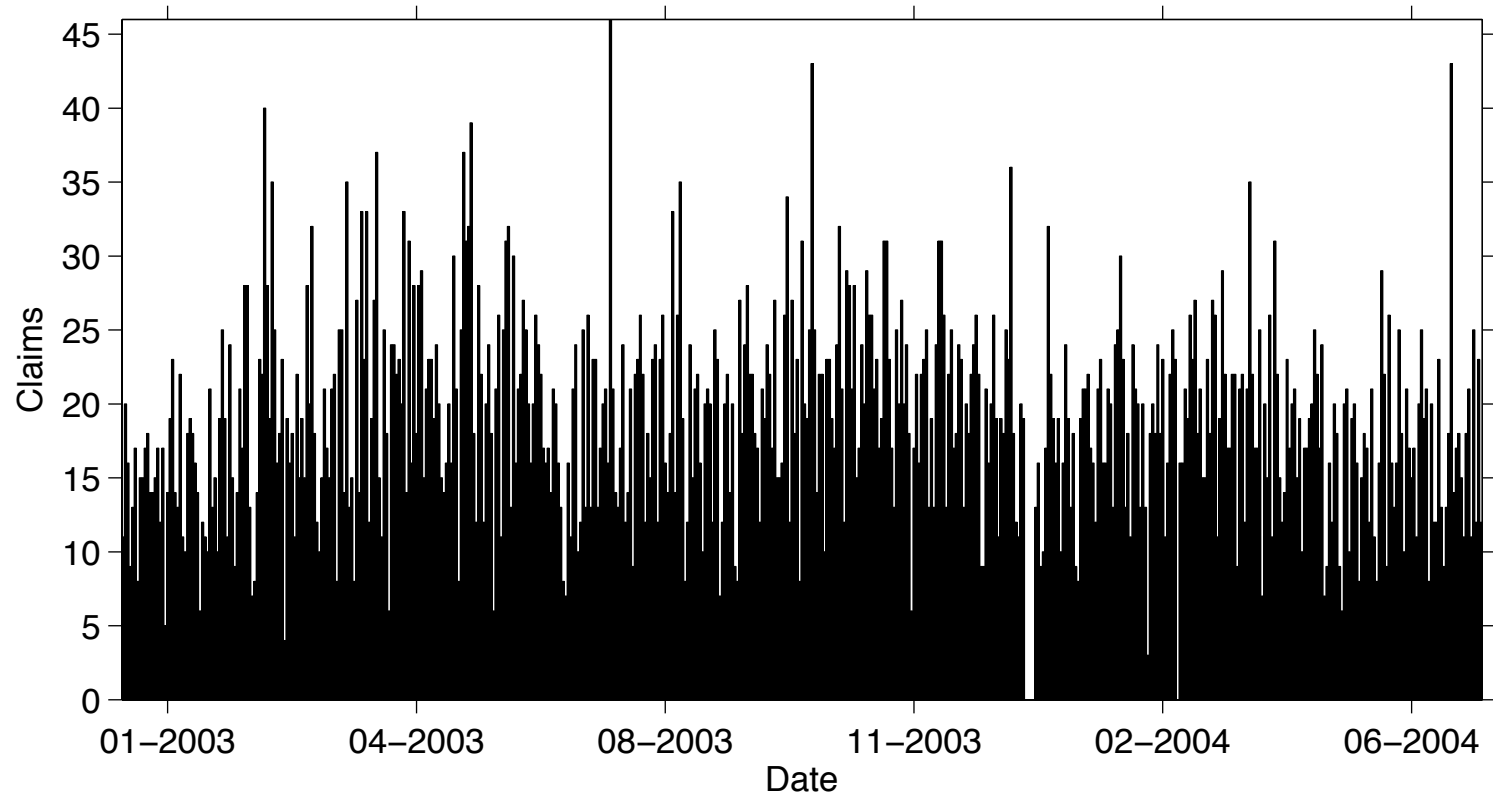


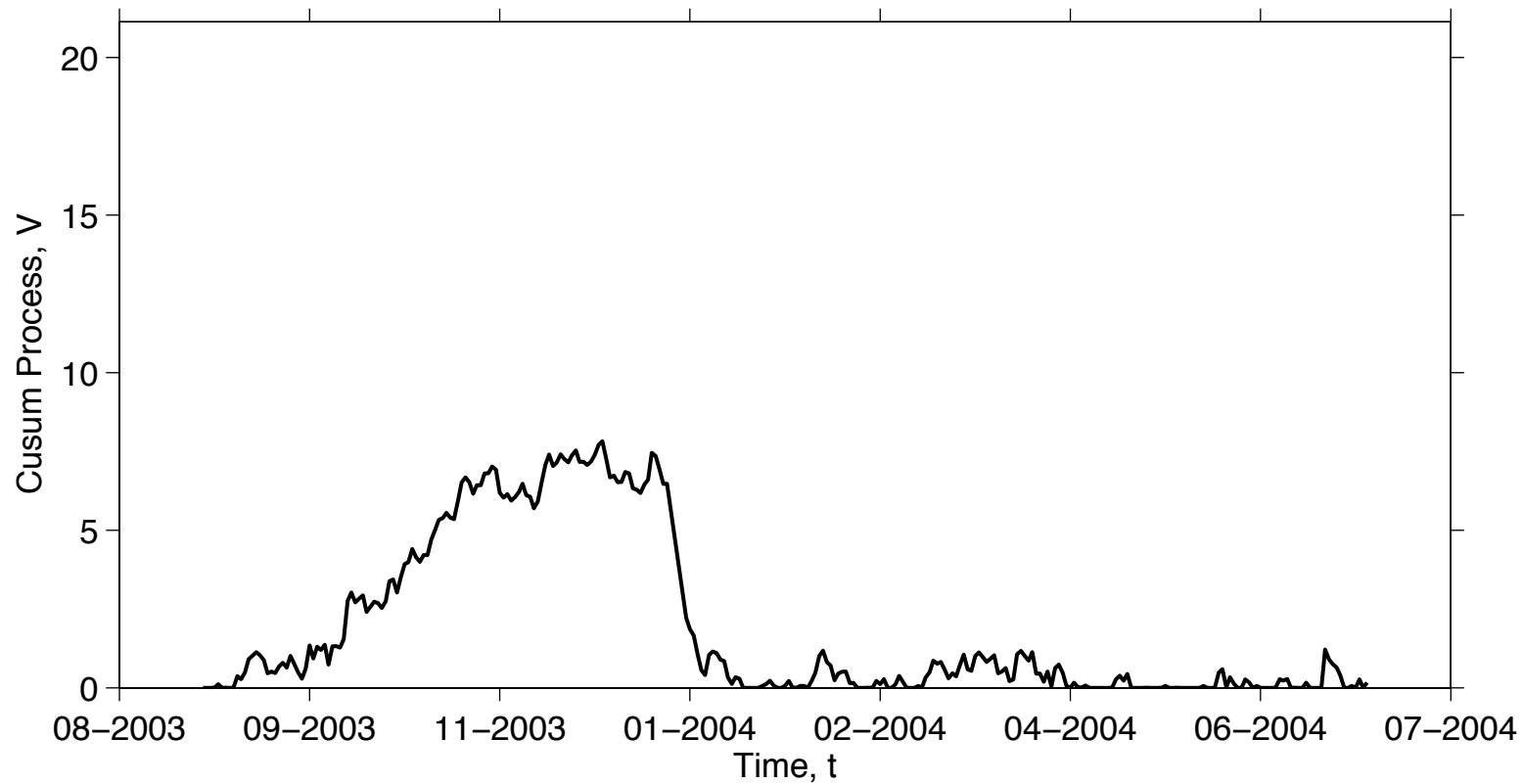
Pour une fois, Verts >> Rouges et bleus











Business-motivated research question

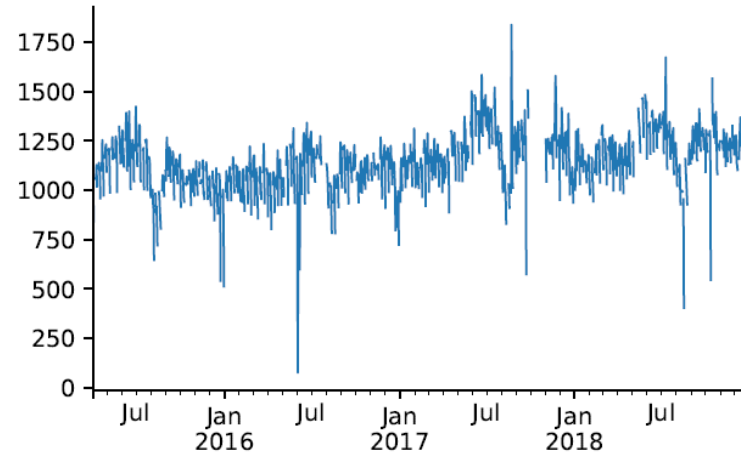
- Can we detect a surge in calls at



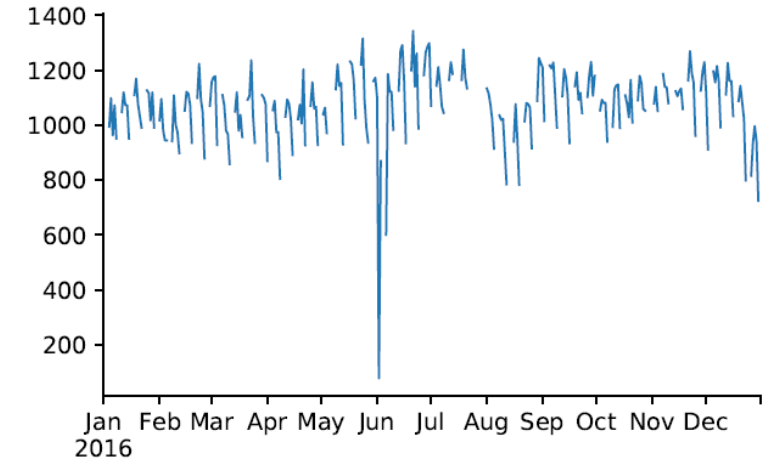
fast enough to onboard extra staff members to handle it?

-> Handle Seasonality

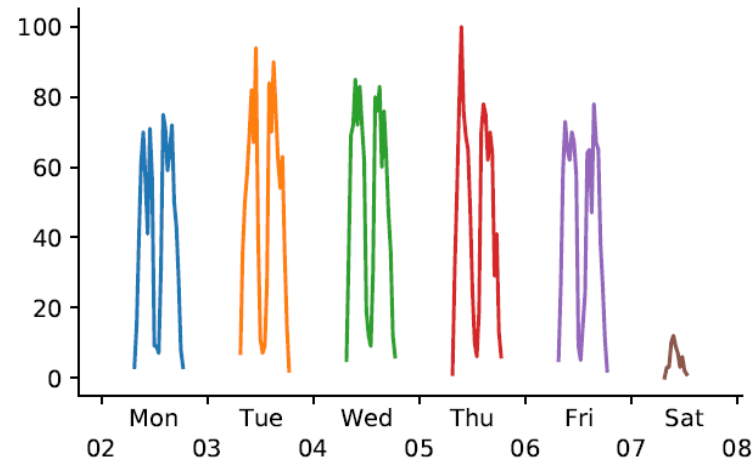
Seasonality



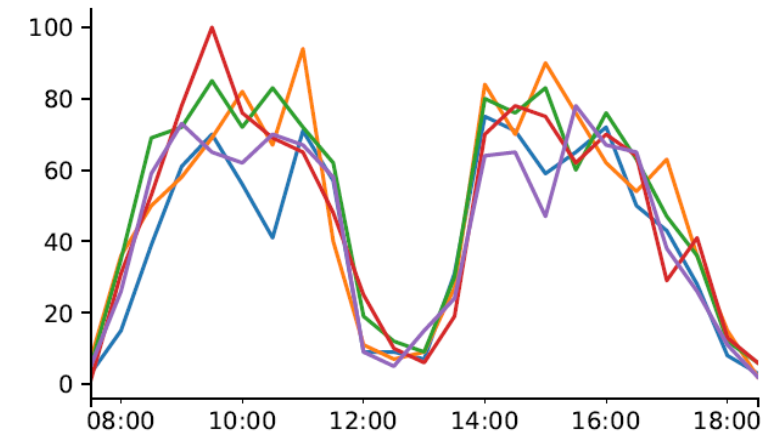
(a) Number of calls each weekday over the period ranging from April 2015 to January 2019.



(b) Number of calls each weekday over the year 2016.



(c) Number of calls during the week ranging from Monday 2nd to Saturday 7th of January 2017.



(d) Number of calls during each day from Monday 2nd to Friday 6th of January 2017.



Figure 1: Seasonality in call arrivals: (a) yearly, (b) monthly, (c) daily and (d) hourly.

What do we want to detect?

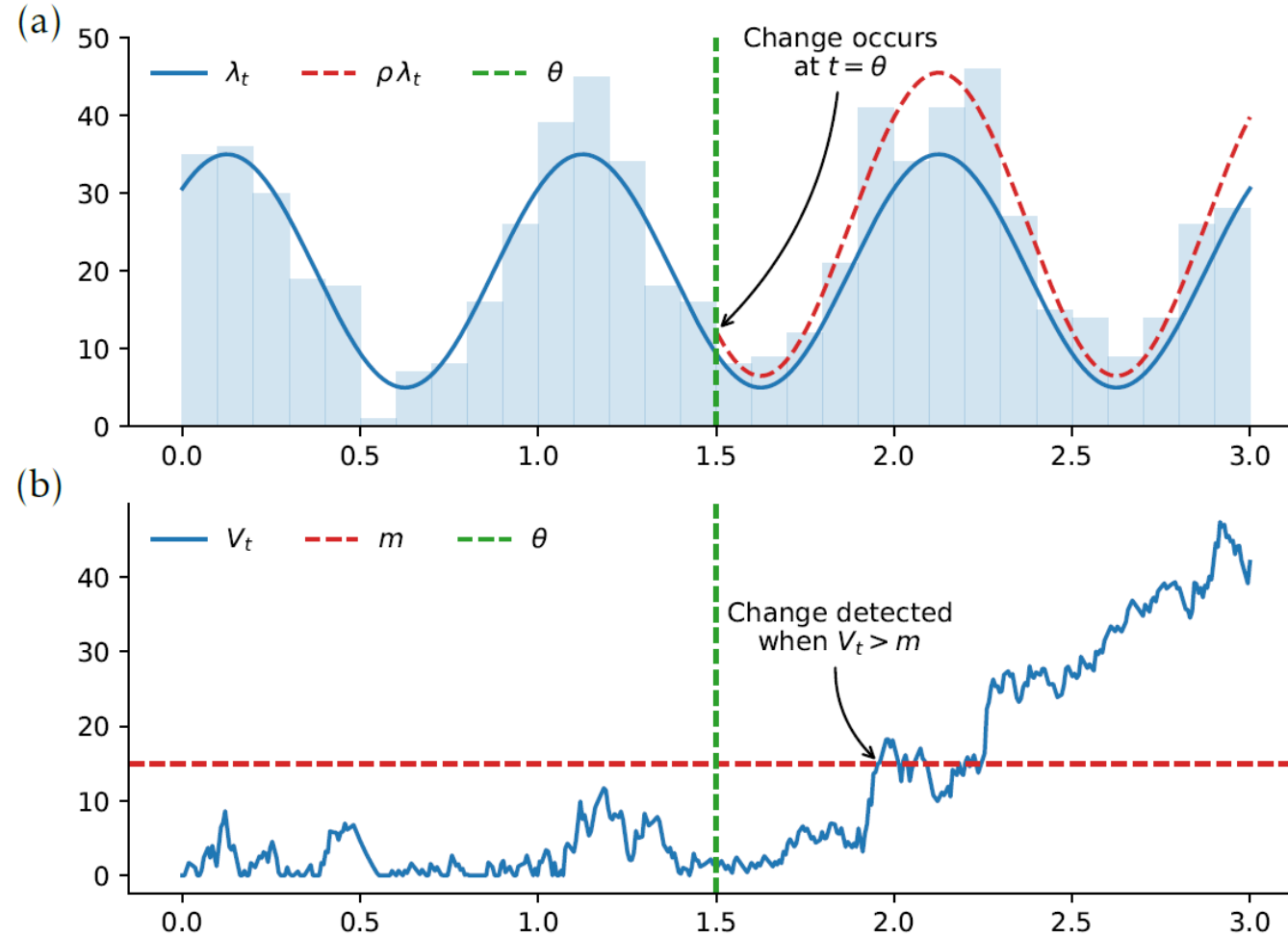


Figure 2: (a) Example of seasonal intensity (solid line) with simulated records and a change-point at time $\theta = 1.5$ with 30% increase of the intensity (dashed line). (b) Sample path of the CUSUM process V .

If we ignore seasonality: false alarms +++

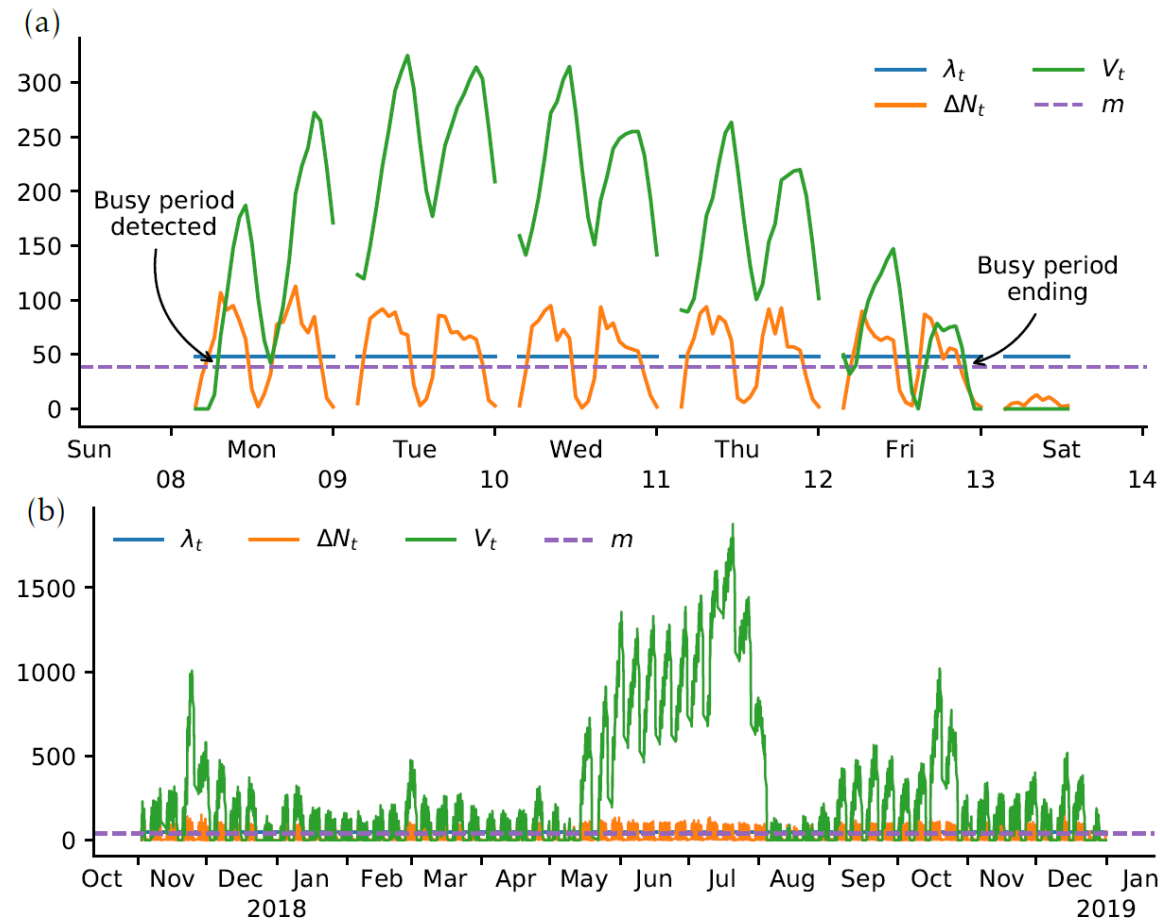


Figure 8: The results of CUSUM with a naive $\lambda_t \equiv \lambda$ over (a) the week starting 2018-01-08 and (b) the whole test period. The number of calls arriving in each half-hour time slot is denoted ΔN_t .

Taking into account seasonality: it « works »!

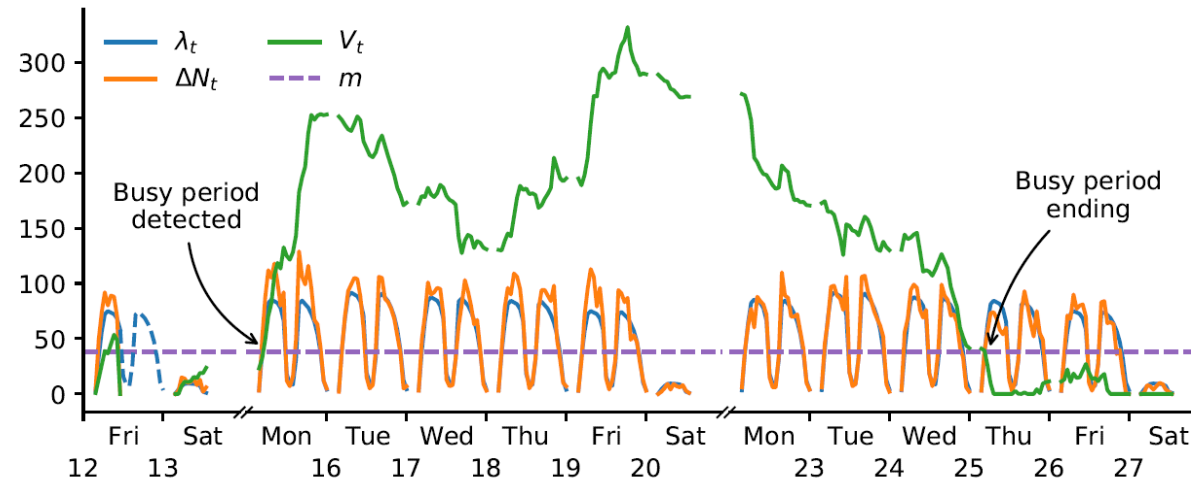


Figure 9: CUSUM algorithm results for a few weeks starting from 2018-10-12. The call center closed at noon on the Friday the 12th, which is unusual. In the CUSUM we modified λ_t to be zero for that afternoon, though here we plotted the original λ_t .

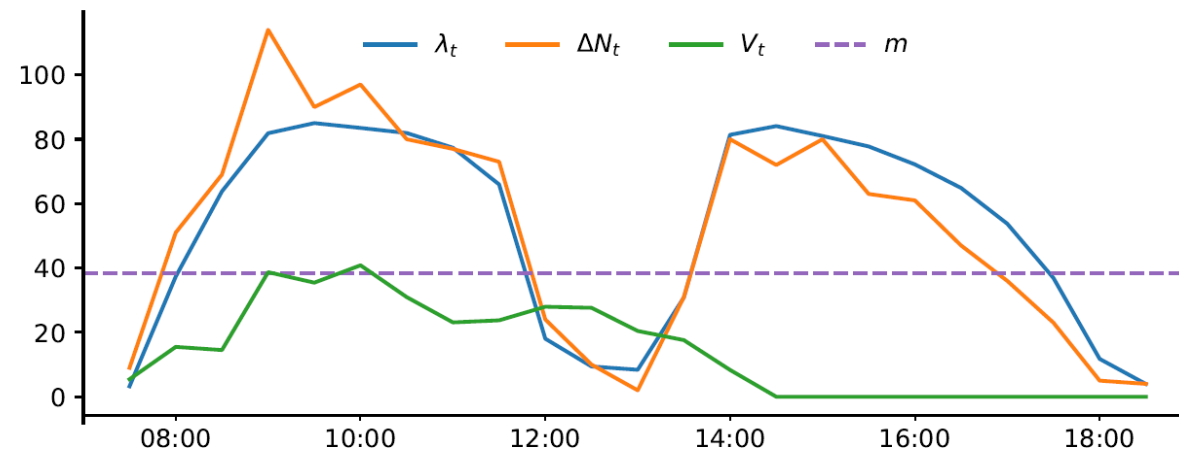


Figure 10: CUSUM algorithm results for the date 2018-06-15.

What if some calls are postponed?

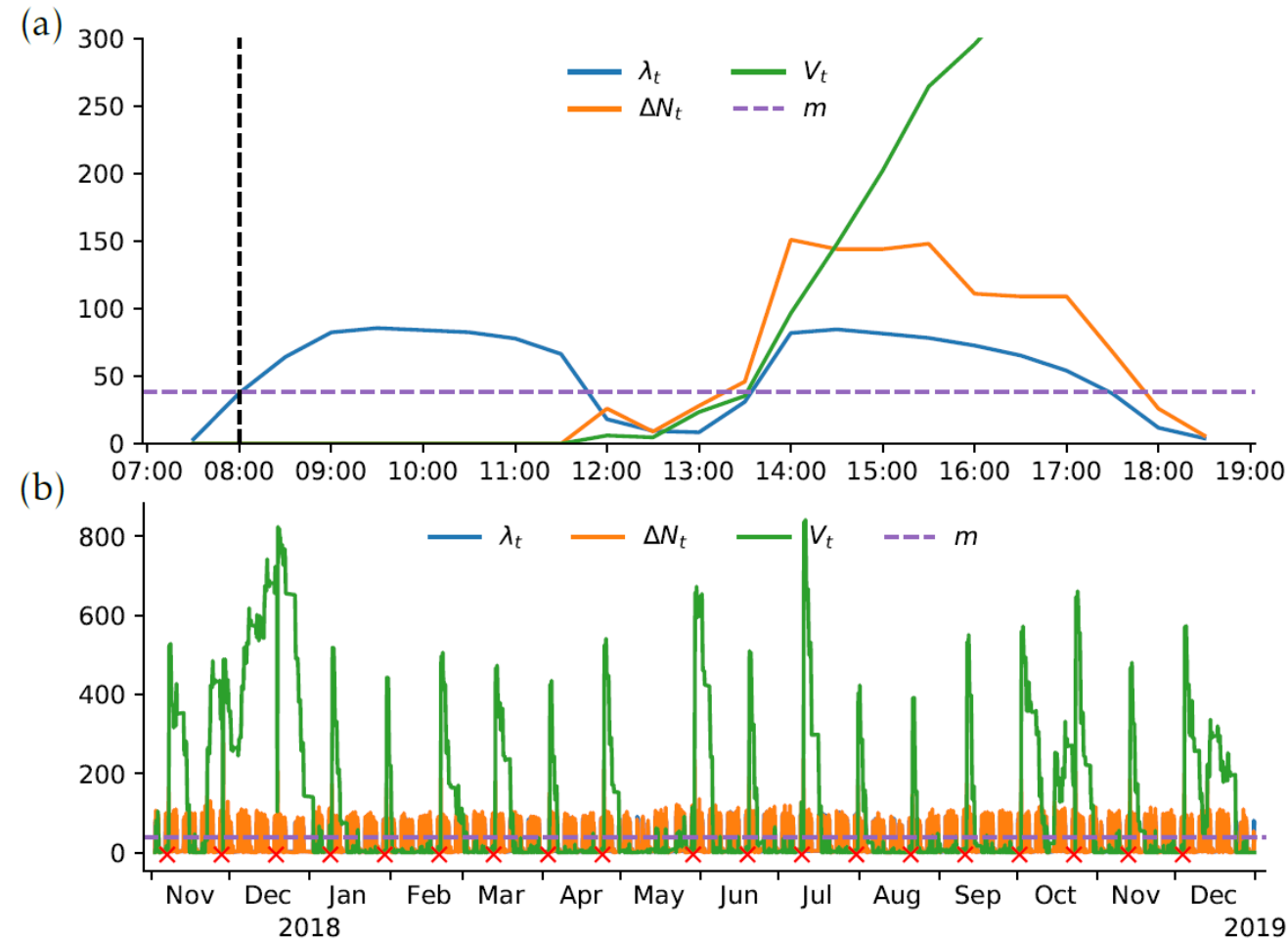


Figure 11: The CUSUM results over (a) 2018-04-03 and (b) the whole test set, where every third Tuesday (marked with a red cross) has had the morning calls shifted into the afternoon.

What if some calls are postponed? Still works!

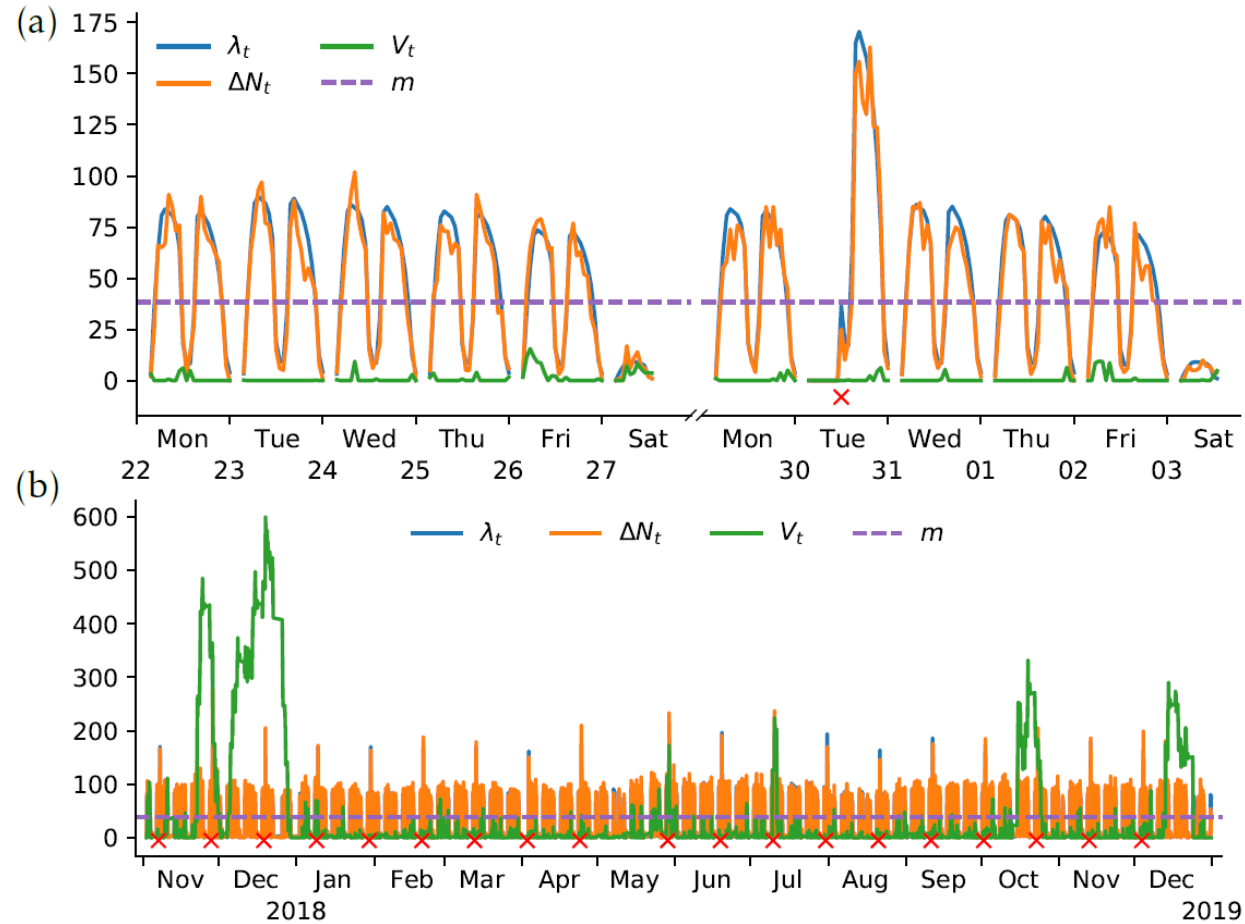


Figure 12: The refitted CUSUM results over (a) a fortnight from 2018-01-22 and (b) the whole test set, where every third Tuesday (marked with a red cross) has had the morning calls shifted into the afternoon.

Potential next steps with



- More detailed presentation with Icare team
- Specific detection for some garage / type of car / region
- Multiple cusums
- Combine with predictive modeling

Wrap-up

- Even if it is not easy to show its optimality,
- Cusum is simple to implement and easy to visualize
- It is faster than anything else (in particular moving window method)
- No miracle!